## Large spread does not imply Benford's Law

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## Abstract

Sharp universal bounds are given for the distance between normalised Lebesgue measure on  $\mathbb{R}/\mathbb{Z}$  and the distribution of log X mod 1, where X is uniform. The results dispel the popular belief that a random variable obeys Benford's Law (at least approximately) whenever its spread is large.

For every real number x, the largest integer not larger than x will be denoted by  $\lfloor x \rfloor$ , and  $\llbracket x \rrbracket := x - \lfloor x \rfloor$  is the fractional (or non-integer) part of x. The base  $\gamma$  logarithm ( $\gamma > 1$ ) of x > 0 is  $\log_{\gamma} x$ ; if used without a subscript, log symbolises the natural logarithm. The sets of natural, non-negative integer, integer, positive real and real numbers are  $\mathbb{N}$ ,  $\mathbb{N}_0$ ,  $\mathbb{Z}$ ,  $\mathbb{R}^+$  and  $\mathbb{R}$ , respectively.

Given any probability measure  $\mu$  on  $\mathbb{R}$ , denote by  $F_{\mu}$  its distribution function, that is,  $F_{\mu}(x) = \mu([-\infty, x])$ . For every measurable map  $T : \mathbb{R} \to \mathbb{R}$  the probability measure  $T\mu$  is defined as  $T\mu(B) := \mu(T^{-1}(B))$  for all Borel sets B. Specifically,  $\llbracket \mu \rrbracket$  is, for every  $\mu$ , concentrated on [0, 1]. The uniform distribution on [a, b] with a < b is denoted by  $U_{a,b}$ . Thus

$$F_{U_{a,b}}(x) = \begin{cases} 0 & \text{if } x < a \,, \\ \frac{x-a}{b-a} & \text{if } a \le x < b \,, \\ 1 & \text{if } x \ge b \,. \end{cases}$$

Given any two probability measures  $\mu, \nu$  on  $\mathbb{R}$ , their Kolmogorov–Smirnov distance  $d_{\infty}(\mu, \nu)$  is

$$d_{\infty}(\mu,\nu) = \sup_{x \in \mathbb{R}} \left| F_{\mu}(x) - F_{\nu}(x) \right|;$$

see e.g. [5] for some details on this metric. Recall that  $\mu$  with  $\mu(\mathbb{R}^+) = 1$ , or a real random variable with distribution  $\mu$ , satisfies *Benford's Law* base  $\gamma$  if and only if  $\log_{\gamma} \mu$  is uniform modulo one [2], i.e., if  $d_{\infty}(\llbracket \log_{\gamma} \mu \rrbracket, U_{0,1})$  equals zero. Contrary to what [4, p.63] may suggest, the uniform distribution  $U_{a,b}$  with  $a \geq 0$  does not even approximately satisfy Benford's Law for any base  $\gamma$ , no matter how large b - a is.

**Theorem 1.** For all  $\gamma > 1$  and  $0 \le a < b$ ,

$$d_{\infty}\left(\llbracket \log_{\gamma} U_{a,b} \rrbracket, U_{0,1}\right) \ge C_{\gamma} > 0, \qquad (1)$$

where

$$C_{\gamma} = \frac{1 - \gamma + \log \gamma + (\gamma - 1)\log(\gamma - 1) - (\gamma - 1)\log\log\gamma}{2(\gamma - 1)\log\gamma}$$

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*Proof.* To verify (1), assume first that a = 0, and let  $l := \lfloor \log_{\gamma} b \rfloor$  and  $\delta := \llbracket \log_{\gamma} b \rrbracket$ . Thus  $b = \gamma^{l+\delta}$ , and

$$F_{\log_{\gamma} U_{0,b}}(x) = \begin{cases} \frac{\gamma^x}{b} & \text{if } x < l + \delta, \\ 1 & \text{if } x \ge l + \delta, \end{cases}$$

from which it follows that, for all  $0 \le x \le 1$ ,

$$F_{\llbracket \log_{\gamma} U_{0,b} \rrbracket}(x) = \sum_{k \in \mathbb{Z}} \left( F_{\log_{\gamma} U_{0,b}}(x+k) - F_{\log_{\gamma} U_{0,b}}(k) \right) = \begin{cases} \frac{\gamma^{1+x-\delta} - \gamma^{1-\delta}}{\gamma - 1} & \text{if } 0 \le x < \delta , \\ 1 - \frac{\gamma^{1-\delta} - \gamma^{x-\delta}}{\gamma - 1} & \text{if } \delta \le x \le 1 . \end{cases}$$

Notice that  $F_{[\log_{\gamma} U_{0,b}]}$  does not depend on l. For the sake of brevity, let  $f_{\delta}(x) := F_{[\log_{\gamma} U_{0,b}]}(x)$  and  $g_{\delta}(x) := f_{\delta}(x) - x$  for all x and  $\delta$ . Note that  $f_{\delta}$  is convex on  $[0, \delta]$  and on  $[\delta, 1]$ , and  $f_0(x) = \frac{\gamma^x - 1}{\gamma - 1}$ , whereas for  $0 < \delta < 1$ 

$$f_{\delta}'(0+) = f_{\delta}'(1-) = \frac{\gamma^{1-\delta}\log\gamma}{\gamma-1}$$

In the latter case, there exists a unique  $0 < \delta^* < 1$  such that  $f'_{\delta^*}(0+) = f'_{\delta^*}(1-) = 1$ ; explicitly  $\gamma^{1-\delta^*} = (\gamma - 1)/\log \gamma$ , and thus

$$\delta^* = \frac{\log \gamma - \log(\gamma - 1) + \log \log \gamma}{\log \gamma}$$

Consequently, for  $0 < \delta < 1$  the graph of  $f_{\delta}$  can have three qualitatively different forms.



Note that  $g_{\delta}$  always attains its maximum at  $x = \delta$ . This suggests introducing the auxiliary function  $\psi$  according to

$$\psi(\delta) := g_{\delta}(\delta) = f_{\delta}(\delta) - \delta = \frac{\gamma - \gamma^{1-\delta}}{\gamma - 1} - \delta.$$

It follows from

$$\psi'(\delta) = \frac{\gamma^{1-\delta}\log\gamma}{\gamma-1} - 1 = \gamma^{\delta^*-\delta} - 1,$$

that  $\psi$  is concave, with  $\psi(0) = \psi(1) = 0$  and  $\psi'(\delta^*) = 0$ . Hence  $\psi(\delta) > 0$  for all  $0 < \delta < 1$ , and

$$\max_{0 \le \delta \le 1} \psi(\delta) = \psi(\delta^*) = \frac{\gamma - \gamma^{1 - \delta^*}}{\gamma - 1} - \delta^* = \frac{\gamma \log \gamma - \gamma + 1}{(\gamma - 1) \log \gamma} - \delta^* = 2C_{\gamma}$$

Assume first that  $0 \leq \delta \leq \delta^*$ . In this case, the function  $g_{\delta}$  has a non-positive minimum at  $x = 1 + \delta - \delta^* > \delta$ , with

$$g_{\delta}(1+\delta-\delta^{*}) = 1 - \frac{\gamma^{1-\delta}-\gamma^{1-\delta^{*}}}{\gamma-1} - 1 - \delta + \delta^{*} = \psi(\delta) - \psi(\delta^{*}),$$

showing that  $\max_{0 \le x \le 1} g_{\delta}(x) = \psi(\delta)$  as well as  $-\min_{0 \le x \le 1} g_{\delta}(x) = \psi(\delta^*) - \psi(\delta)$ . Similarly, if  $\delta^* < \delta < 1$  then  $g_{\delta}$  has a negative minimum at  $x = \delta - \delta^* < \delta$ , with

$$g_{\delta}(\delta - \delta^*) = \frac{\gamma^{1 - \delta^*} - \gamma^{1 - \delta}}{\gamma - 1} - \delta + \delta^* = \psi(\delta) - \psi(\delta^*).$$

For all  $0 \leq \delta < 1$ , therefore,

$$\max_{0 \le x \le 1} g_{\delta}(x) = \psi(\delta), \quad -\min_{0 \le x \le 1} g_{\delta}(x) = \psi(\delta^*) - \psi(\delta),$$

and consequently

$$\max_{0 \le x \le 1} |g_{\delta}(x)| = \max \left\{ \psi(\delta), \psi(\delta^*) - \psi(\delta) \right\} \ge \frac{1}{2} \psi(\delta^*) = C_{\gamma},$$

which establishes (1) for the case a = 0.

To verify (1) for a > 0 assume for the time being that  $\log_{\gamma} a = k \in \mathbb{Z}$ , and let  $l := \lfloor \log_{\gamma} b \rfloor$  and  $\delta := \llbracket \log_{\gamma} b \rrbracket$  as before; for convenience set  $m := l - k \in \mathbb{N}_0$ . A short computation confirms that

$$f_{m,\delta}(x) := F_{\llbracket \log_{\gamma} U_{a,b} \rrbracket}(x) = \begin{cases} \frac{\gamma^x - 1}{\gamma - 1} \cdot \frac{\gamma^{m+1} - 1}{\gamma^{m+\delta} - 1} & \text{if } 0 \le x < \delta \,, \\ 1 - \frac{\gamma - \gamma^x}{\gamma - 1} \cdot \frac{\gamma^m - 1}{\gamma^{m+\delta} - 1} & \text{if } \delta \le x \le 1 \,. \end{cases}$$

Notice that  $f_{m,\delta} \to f_{\delta}$  uniformly on [0,1] as  $m \to \infty$ . Let again  $g_{m,\delta}(x) := f_{m,\delta}(x) - x$  and observe that

$$g_{m,\delta}(x) - g_{\delta}(x) = f_{m,\delta}(x) - f_{\delta}(x) = \Delta_{m,\delta}(x) ,$$

where  $\Delta_{m,\delta}$  is given by

$$\Delta_{m,\delta}(x) = \begin{cases} \frac{\gamma^x - 1}{\gamma - 1} \cdot \frac{\gamma^{1-\delta} - 1}{\gamma^{m+\delta} - 1} & \text{if } 0 \le x < \delta \,, \\ \frac{\gamma^{1-\delta} - \gamma^{x-\delta}}{\gamma - 1} \cdot \frac{\gamma^{\delta} - 1}{\gamma^{m+\delta} - 1} & \text{if } \delta \le x \le 1 \,. \end{cases}$$

Obviously,  $\Delta_{m,\delta} \geq 0$  with  $\Delta_{m,\delta}(0) = \Delta_{m,\delta}(1) = 0$ , and  $\Delta_{m,0} = 0$  for all  $m \geq 1$ . Furthermore, for  $0 < \delta < 1$  the function  $\Delta_{m,\delta}$  is convex and increasing on  $[0, \delta]$ , and concave and decreasing on  $[\delta, 1]$ . Since both  $g_{\delta}$  and  $\Delta_{m,\delta}$  attain their respective maximal value at  $x = \delta$ ,

$$\max_{0 \le x \le 1} g_{m,\delta}(x) = g_{m,\delta}(\delta) = g_{\delta}(\delta) + \Delta_{m,\delta}(\delta) = \psi(\delta) + \Delta_{m,\delta}(\delta).$$

If  $0 \le \delta \le \delta^*$  then, with the appropriate  $0 \le \xi \le 1$ ,

$$\begin{aligned} \max_{0 \le x \le 1} g_{m,\delta}(x) - \min_{0 \le x \le 1} g_{m,\delta}(x) &= g_{m,\delta}(\delta) - g_{m,\delta}(\xi) \\ &\ge g_{m,\delta}(\delta) - g_{m,\delta}(1 + \delta - \delta^*) \\ &= \psi(\delta) + \Delta_{m,\delta}(\delta) - g_{\delta}(1 + \delta - \delta^*) - \Delta_{m,\delta}(1 + \delta - \delta^*) \\ &= \psi(\delta^*) + \Delta_{m,\delta}(\delta) - \Delta_{m,\delta}(1 + \delta - \delta^*) \\ &\ge \psi(\delta^*) \\ &= 2C_{\gamma} \,. \end{aligned}$$

The same argument applies for  $\delta^* < \delta < 1$  with  $1 + \delta - \delta^*$  replaced by  $\delta - \delta^*$ . Thus

$$\max_{0 \le x \le 1} g_{m,\delta}(x) - \min_{0 \le x \le 1} g_{m,\delta}(x) \ge 2C_{\gamma}$$

$$\tag{2}$$

holds for all  $m \in \mathbb{N}_0$  and  $0 \leq \delta < 1$ , and this in turn implies (1) since

$$\max_{0 \le x \le 1} |g_{m,\delta}(x)| \ge \frac{1}{2} \left( \max_{0 \le x \le 1} g_{m,\delta}(x) - \min_{0 \le x \le 1} g_{m,\delta}(x) \right) \ge C_{\gamma}.$$

Overall, therefore, the proof is complete if a = 0 or  $\log_{\gamma} a \in \mathbb{Z}$ .

Finally, assume that a > 0 does not satisfy  $\log_{\gamma} a \in \mathbb{Z}$ , that is, the number  $\tau := [\log_{\gamma} a]$  lies strictly between 0 and 1. Note that

$$\llbracket \log_{\gamma} U_{a,b} \rrbracket = \llbracket \log_{\gamma} U_{a\gamma^{-\tau}, b\gamma^{-\tau}} + \tau \rrbracket$$

and clearly  $\log_{\gamma}(a\gamma^{-\tau}) \in \mathbb{Z}$ . It is readily verified that, for every non-atomic probability measure  $\mu$  on  $\mathbb{R}$  and every  $t \in \mathbb{R}$ ,

$$F_{\llbracket \mu + t \rrbracket}(x) = \begin{cases} F_{\llbracket \mu \rrbracket}(x + 1 - \llbracket t \rrbracket) - F_{\llbracket \mu \rrbracket}(1 - \llbracket t \rrbracket) & \text{if } 0 \le x < \llbracket t \rrbracket, \\ F_{\llbracket \mu \rrbracket}(x - \llbracket t \rrbracket) + 1 - F_{\llbracket \mu \rrbracket}(1 - \llbracket t \rrbracket) & \text{if } \llbracket t \rrbracket \le x \le 1, \end{cases}$$

and therefore also

$$G_{\llbracket \mu + t \rrbracket}(x) = \begin{cases} G_{\llbracket \mu \rrbracket}(x + 1 - \llbracket t \rrbracket) - G_{\llbracket \mu \rrbracket}(1 - \llbracket t \rrbracket) & \text{if } 0 \le x < \llbracket t \rrbracket, \\ G_{\llbracket \mu \rrbracket}(x - \llbracket t \rrbracket) - G_{\llbracket \mu \rrbracket}(1 - \llbracket t \rrbracket) & \text{if } \llbracket t \rrbracket \le x \le 1, \end{cases}$$

where generally  $G_{\mu}(x) := F_{\mu}(x) - x$ . In particular,

$$\max_{0 \le x \le 1} G_{\llbracket \mu + t \rrbracket}(x) - \min_{0 \le x \le 1} G_{\llbracket \mu + t \rrbracket}(x) = \max_{0 \le x \le 1} G_{\llbracket \mu \rrbracket}(x) - \min_{0 \le x \le 1} G_{\llbracket \mu \rrbracket}(x), \quad (3)$$

which merely expresses the intuitively obvious fact that the *span* (i.e. the difference between maximal and minimal value) of  $G_{\llbracket \mu \rrbracket}$  is not affected by the rotation caused by adding (modulo one) any number t. With the notation introduced earlier,  $G_{\llbracket \log_{\gamma} U_{a\gamma^{-\tau},b\gamma^{-\tau}}\rrbracket} = g_{m,\delta}$ , where  $m = \lfloor \log_{\gamma} b/a \rfloor$  and  $\delta = \llbracket \log_{\gamma} b/a \rrbracket$ . Combining (2) and (3) for  $\mu = \log_{\gamma} U_{a\gamma^{-\tau},b\gamma^{-\tau}}$  and  $t = \tau$  therefore yields

$$\max_{0 \le x \le 1} G_{\llbracket \log_{\gamma} U_{a,b} \rrbracket}(x) - \min_{0 \le x \le 1} G_{\llbracket \log_{\gamma} U_{a,b} \rrbracket}(x) \ge 2C_{\gamma}.$$

This completes the proof.

**Remark 2.** (i) As the above argument shows, the constant  $C_{\gamma}$  in (1) is best possible: For every  $C > C_{\gamma}$  there exist a, b with 0 < a < b such that  $d_{\infty}([\log_{\gamma} U_{a,b}]], U_{0,1}) < C$ .

(ii) It follows from the first part of the proof of Theorem 1 that, for all  $\gamma > 1$  and b > 0,

$$d_{\infty}\left(\llbracket \log_{\gamma} U_{0,b} \rrbracket, U_{0,1}\right) = \Psi(\log_{\gamma} b),$$

with the continuous, 1-periodic function  $\Psi: x \mapsto \max\{\psi(\llbracket x \rrbracket), 2C_{\gamma} - \psi(\llbracket x \rrbracket)\} = C_{\gamma} + |\psi(\llbracket x \rrbracket) - C_{\gamma}|.$ 

(iii) Note that  $\gamma \mapsto C_{\gamma}$  is monotonically increasing, with  $\lim_{\gamma \to 1^+} C_{\gamma} = 0$  and  $\lim_{\gamma \to \infty} C_{\gamma} = \frac{1}{2}$ . For  $\gamma = 10$ , the most important special case in view of Benford's Law, one finds  $C_{10} \approx 0.13442$ .

(iv) Satisfactory though it may be, Theorem 1 has a small shortcoming: For every probability measure  $\mu$  on  $\mathbb{R}$ , the measure  $\llbracket \mu \rrbracket$  naturally lives on  $\mathbb{T} = \mathbb{R}/\mathbb{Z}$  rather than on [0, 1], but  $d_{\infty}$  is unsuitable for measures on  $\mathbb{T}$ . Specifically,  $\mathbb{T}$  is a compact metric space when endowed with the metric  $d(x + \mathbb{Z}, y + \mathbb{Z}) := \min_{k \in \mathbb{Z}} |x - y + k|$ , and consequently  $\mathcal{P}(\mathbb{T})$ , the space of all probability measures on  $\mathbb{T}$  with the topology of weak convergence, is compact and metrizable [3]. A natural metric inducing this topology is the Kantorovich–Wasserstein distance  $d_K$  defined as

$$d_{K}(\mu,\nu) := \sup \left\{ \left| \int_{\mathbb{T}} f \, \mathrm{d}\mu - \int_{\mathbb{T}} f \, \mathrm{d}\nu \right| : f \in C_{\mathbb{R}}(\mathbb{T}), \operatorname{Lip} f \leq 1 \right\}.$$

Unlike  $d_{\infty}$ , the metric  $d_K$  on  $\mathcal{P}(\mathbb{T})$  is invariant under isometries of  $\mathbb{T}$ : If  $T : \mathbb{T} \to \mathbb{T}$  is any isometry, then  $d_K(T\mu, T\nu) = d_K(\mu, \nu)$  for all  $\mu, \nu \in \mathcal{P}(\mathbb{T})$ . Explicit practicable formulae for  $d_K$  have been derived in [1]. A truly satisfactory variant of Theorem 1, therefore, would consider  $[\log_{\gamma} U_{a,b}]$  an element of  $\mathcal{P}(\mathbb{T})$  and provide a lower bound for its distance from  $\lambda_{\mathbb{T}}$ , the uniform distribution on  $\mathbb{T}$ . Such a result can indeed be achieved using parts of the proof of Theorem 1 even though the necessary calculations are significantly more involved. The final result, however, is even slightly simpler than (1): For all  $\gamma > 1$  and  $0 \le a < b$ ,

$$d_K\left(\llbracket \log_{\gamma} U_{a,b} \rrbracket, \lambda_{\mathbb{T}}\right) \ge \log_{\gamma} \frac{1+\sqrt{\gamma}}{2} - \frac{1}{4} =: \frac{1}{4} \Phi\left(\frac{1}{4}\log\gamma\right) > 0, \qquad (4)$$

where  $\Phi$  is the real-analytic odd function  $\Phi(x) = x^{-1} \log \cosh x$ ; as in the case of (1), the inequality (4) is best possible in the sense of (i).

## References

- C.A. Cabrelli and U.M. Molter, The Kantorovich metric for probability measures on the circle, J. Comput. Appl. Math. 57 (1995), 345–361.
- [2] P. Diaconis, The distribution of leading digits and uniform distribution mod 1, Ann. Probab. 5 (1979), 72–81.
- [3] R.M. Dudley, *Real analysis and probability*, Revised reprint of the 1989 original, Cambridge University Press (2002).
- [4] W. Feller, An introduction to probability theory and its applications, Vol. II. 2nd ed., John Wiley and Sons, New York (1971).
- [5] A.L. Gibbs and F.E. Su, On choosing and bounding probability metrics, Int. Stat. Rev. 70 (2002), 419–435.